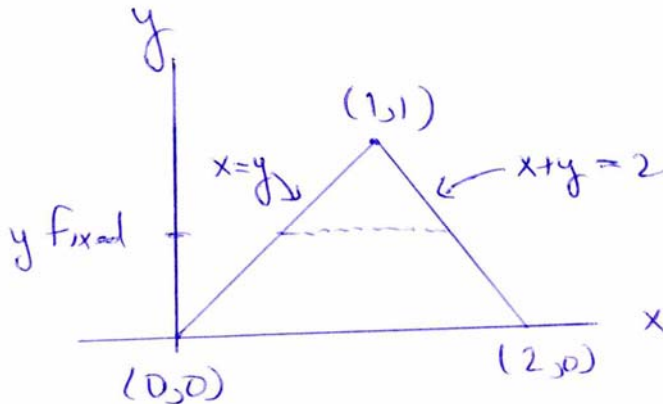


Stat 334 Midterm Solution

1. Suppose X and Y have joint pdf $f(x, y) = 1$ with support defined by the triangle with vertices $(0,0)$, $(2,0)$ and $(1,1)$. Find
- (3 marks) the marginal pdf of Y .

The support of the joint distribution is shown in the sketch.



To find the marginal pdf of Y , we hold Y fixed and integrate over the possible values of X . From the sketch we have

$$f_2(y) = \int_{x=y}^{x=2-y} 1 dx = 2 - 2y, \quad 0 < y < 1$$

You must show the marginal support of Y that does not depend on X .

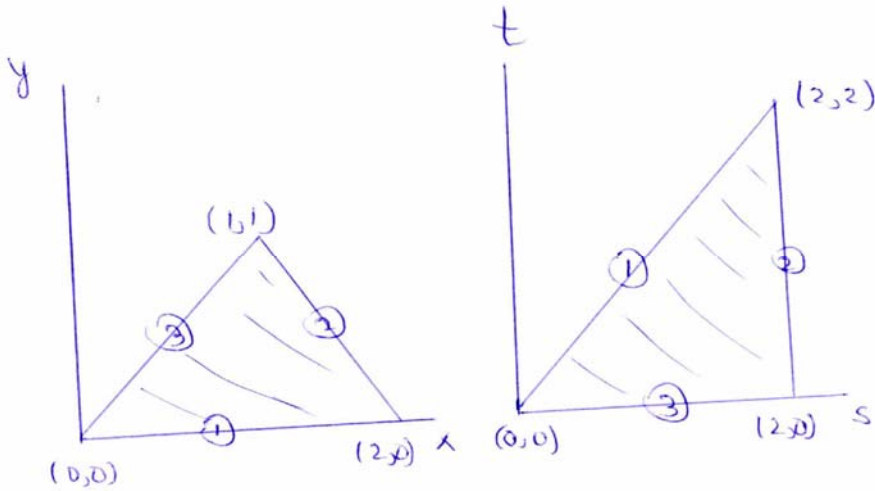
- (3 marks) $E[XY]$

Note from the sketch it is easier to integrate first with respect to x .

$$E[XY] = \int_{y=0}^1 \int_{x=y}^{2-y} xy dx dy = \int_{y=0}^1 y \left[\frac{(2-y)^2}{2} - \frac{y^2}{2} \right] dy = 2 \int_{y=0}^1 (y - y^2) dy = 1/3$$

- (4 marks) the joint pdf of $S = X + Y, T = X - Y$

First we look at the support of S and T . From the sketch and mapping edges onto edges, we see that the support is the triangle $0 < t < s < 2$.



- 1: $y = 0, 0 < x < 2 \Rightarrow t = s, 0 < s < 2$
- 2: $x + y = 2, 0 < y < 1 \Rightarrow s = 2, 0 < t < 2$
- 3: $y = x, 0 < x < 1 \Rightarrow t = 0, 0 < s < 2$

First we look at the support of S and T . From the sketch and mapping edges onto edges, we see that the support is the triangle $0 < t < s < 2$.

Next we find the Jacobian. Here

$$J = \left| \det \begin{pmatrix} \frac{\partial s}{\partial x} & \frac{\partial s}{\partial y} \\ \frac{\partial t}{\partial x} & \frac{\partial t}{\partial y} \end{pmatrix} \right| = \left| \det \begin{pmatrix} 1 & 1 \\ 1 & -1 \end{pmatrix} \right| = 2$$

Since $f(x, y) = 1$ is constant, we have the joint pdf of S and T is

$$g(s, t) = \frac{1}{2}, 0 < t < s < 2$$

2. Suppose U is a discrete random variable with probability function

$$p(u) = \left(\frac{1}{2}\right)^u, u = 1, 2, \dots$$

a) (3 marks) Show that the moment generating function of U is

$$m(t) = \frac{e^t}{2 - e^t}, t < \ln(2)$$

We have

$$\begin{aligned}
m(t) &= E[\exp\{Ut\}] \\
&= \sum_{u=1}^{\infty} e^{ut} (1/2)^u \\
&= \frac{e^t}{2} \left(1 + \left(\frac{e^t}{2}\right) + \left(\frac{e^t}{2}\right)^2 + \dots\right) \quad (\text{an infinite geometric series}) \\
&= \frac{e^t/2}{1 - e^t/2}, \quad e^t/2 < 1 \\
&= \frac{e^t}{2 - e^t}, \quad t < \ln(2)
\end{aligned}$$

b) (2 marks) Find $E[U]$.

$$E[U] = M'(0) \text{ and } M'(t) = \frac{(2 - e^t)e^t - e^t(-e^t)}{(2 - e^t)^2} \text{ so } M'(0) = \frac{1+1}{(1)^2} = 2$$

3. (a) (2 marks) For any two random variables X and Y, define $E[X | Y]$.

$$\text{First we } H(y) = E[X | Y = y] = \int_{\substack{\text{all } x \text{ with} \\ y \text{ fixed}}} xf(x|y)dx \text{ and then } E[X | Y] = h(Y)$$

(b) (3 marks). Suppose $U \sim \text{unif}(0,1)$ and, given $U = u$, $Y \sim \text{binomial}(n,u)$. Find $\text{Var}[Y]$

$\text{Var}[Y] = E[\text{Var}[Y | U]] + \text{Var}[E[Y | U]]$ and we have (from simple properties of the binomial distribution)

$$\begin{aligned}
E[Y | U = u] &= nu \\
\text{Var}[Y | U = u] &= nu(1 - u)
\end{aligned}$$

and so

$$\begin{aligned}
E[Y | U] &= nU \\
\text{Var}[Y | U] &= nU(1 - U) = nU - nU^2
\end{aligned}$$

Since $U \sim \text{unif}(0,1)$ we have $f(u) = 1$, $0 < u < 1$ and so

$$E[U] = \int_{u=0}^1 u du = 1/2, \quad E[U^2] = \int_{u=0}^1 u^2 du = 1/3, \quad \text{Var}[U] = 1/3 - (1/2)^2 = 1/12$$

Finally

$$\begin{aligned}
\text{Var}[Y] &= E[\text{Var}[Y | U]] + \text{Var}[E[Y | U]] \\
&= E[nU(1 - U)] + \text{Var}[nU] \\
&= nE[U^2] - nE[U] + n^2\text{Var}[U] \\
&= n/3 - n/2 + n^2/12 \\
&= n/6 + n^2/12
\end{aligned}$$